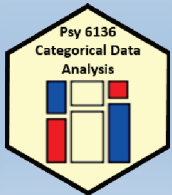




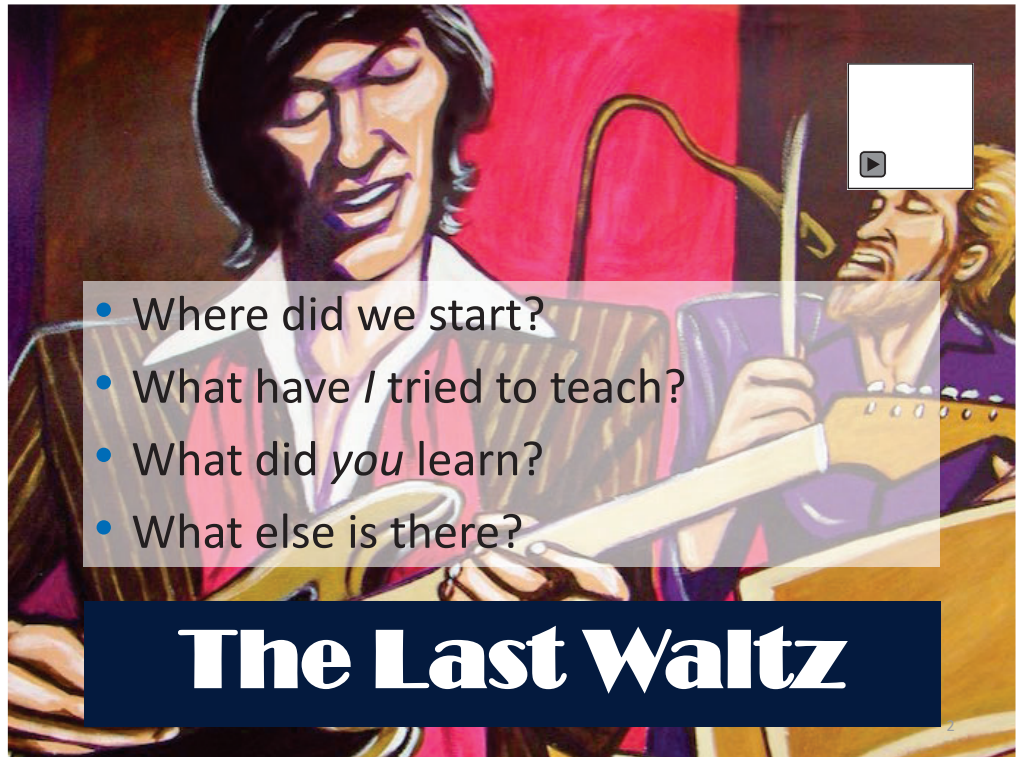
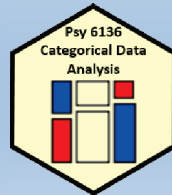
The Last Waltz



Michael Friendly

Psych 6136

<https://friendly.github.io/6136>



- Where did we start?
- What have I tried to teach?
- What did *you* learn?
- What else is there?

The Last Waltz

My goals



Start with descriptive, hypothesis testing methods, then progress to model-based methods



Visual tools for thinking & understanding

Sieve plots, mosaic plots, spineplots, ...
Correspondence analysis: best 2D summary
Effect plots, Data + Model plots



Build from simple, loglinear models to more complex ones

01: Overview

- Categorical data involves some new ideas
 - Discrete variables: `unordered` or `ordered`
 - Counts, frequencies as outcomes
- New / different data structures & functions
 - tables – 1-way, 2-way, 3-way, ... `table()`, `xtabs()`
 - similar in matrices or arrays `matrix()`, `array()`
 - datasets:
 - frequency form
 - case form
- Graphical methods: often use area ~ Freq
 - Consider: graphical comparisons, effect order
- Models: Most are \cong natural extensions of `lm()`

Categorical data: Structures

Categorical (frequency) data appears in various forms

- **Tables:** often the result of `table()` or `xtabs()`

- 1-way
- 2-way — 2×2 , $r \times c$
- 3-way

Gender compared to handedness

	Handed	
	Left	Right
Female	7	46
Male	1	53

margins

- **Matrices:** `matrix()`, with row & col names

- **Arrays:** `array()`, with `dimnames()`

- **Data frames**

- Case form (individual observations)
- Frequency form

	Hair	Eye	Freq
1	Black	Brown	68
2	Brown	Brown	119
3	Red	Brown	14
4	Black	Blue	20
5	Brown	Blue	84
6	Red	Blue	17

Effect ordering: Frequency tables

- Effect ordering and high-lighting for tables

Table: Hair color - Eye color data: Effect ordered

Eye color	Hair color			
	Black	Brown	Red	Blond
Brown	68	119	26	79
Hazel	15	54	14	10
Green	5	29	14	16
Blue	20	84	17	94

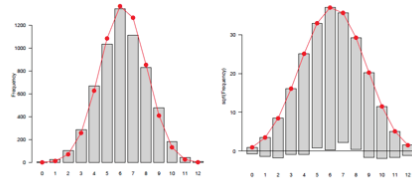
Model: Independence: [Hair][Eye] $\chi^2(9) = 138.29$
 Color coding: <-1 <-2 <-1 0 >1 >2 >4
 n in each cell: n < expected n > expected

The pattern is clearer when the eye colors are **permuted**: light hair goes with light eyes & vice-versa

1-way tables: graphs

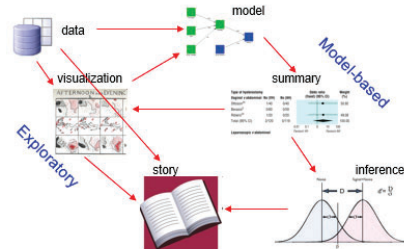
For a particular distribution in mind:

- Plot the data together with the fitted frequencies
- Better still: **hanging rootogram**: freq on sqrt scale; hang bars from fitted values



Data, pictures, models & stories

Now, tell the story!



02: Discrete distributions

- Discrete distributions are the building blocks for categorical data analysis
 - Typically consist of basic counts of occurrences, with varying frequencies
 - Most common: binomial, Poisson, negative binomial
 - Others: geometric, log-series
- Fit with `goodfit()`; plot with `rootogram()`
 - Diagnostic plots: `Ord_plot()`, `distplot()`
- Models with predictors
 - Binomial → logistic regression
 - Poisson → poisson regression; logliner models
 - These are special cases of **generalized** linear models

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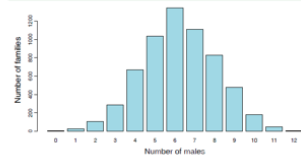
Examples: binomial

Human sex ratio (Geissler, 1889): Is there evidence that $P(\text{male}) = 0.5$?

Saxony families

Saxony families with 12 children having $k = 0, 1, \dots, 12$ sons.

k	0	1	2	3	4	5	6	7	8	9	10	11	12
n_k	3	24	104	286	670	1033	1343	1112	829	478	181	45	7



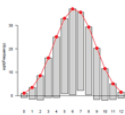
Common discrete distributions

Distribution	Counts, k	Values of X	$P(X=k)$	Mean, $E(X)$	Var, $V(X)$
Bernoulli(p)	Success in 1 trial	$k \in \{0, 1\}$	$p^k(1-p)^{1-k}$	p	$p(1-p)$
Binomial(n, p)	# successes in n trials	$0, 1, \dots, n$	$\binom{n}{k} p^k(1-p)^{n-k}$	np	$np(1-p)$
Geometric(p)	# of trials to 1 st success	$0, 1, 2, \dots$	$p(1-p)^k$	$\frac{1-p}{p}$	$\frac{1-p}{p^2}$
Neg. binomial(k, p)	# of trials to k th success	$0, 1, 2, \dots$	$\binom{n+k-1}{k} p^k(1-p)^n$	$\frac{k(1-p)}{p}$	$\frac{k(1-p)}{p^2}$
Poisson(λ)	# of events in interval	$0, 1, 2, \dots$	$\frac{\lambda^k e^{-\lambda}}{k!}$	λ	λ
Log series(p)	# of types observed	$0, 1, 2, \dots$	$\frac{p^k}{k \log(1-p)}$		

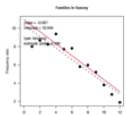
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Graphing discrete distributions

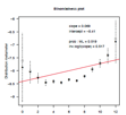
Rootograms



Ord plots



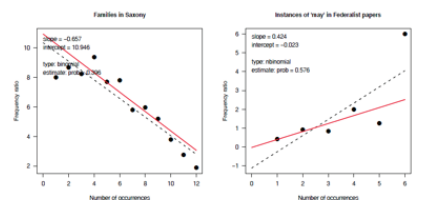
Robust distribution plots



Ord plots: Examples

Ord plots for the Saxony and Federalist data

```
> Ord_plot(Saxony, main = "Families in Saxony", gp=gpar(cex=1, pch=16))
> Ord_plot(Federalist, main = "Instances of 'may' in Federalist papers", gp=gpar(cex=1, pch=16))
```



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03: Two-way tables

- Two-way tables summarize frequencies of two categorical factors
 - 2×2 : a special case, with **odds ratio** as a measure
 - $r \times c$: factors can be **unordered** or **ordered**
 - $r \times c \times k$: stratified tables, $r \times c$ with groups or circumstances
- Tests & measures of association
 - Pearson χ^2 , LR G^2 : **general association**
 - More powerful **CMH tests** for ordered factors
- Visualization
 - 2×2 : fourfold plots
 - $r \times c$: sieve diagrams, tile plots, ...
 - More graphical methods to come ...

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Measures of association

- 2 × 2 tables
- Odds ratio

$$\theta = \frac{\text{odds}(B_1 | A_1)}{\text{odds}(B_1 | A_2)} = \frac{n_{11} / n_{12}}{n_{21} / n_{22}}$$

- Phi coefficient
 - Analog of correlation
 - $\phi^2 = \%$ of variance

$$\phi = \frac{n_{11}n_{22} - n_{12}n_{21}}{n_1 n_2} = \pm \sqrt{\chi^2 / n}$$

- r × c tables

- Cramer's V – generalization of phi

$$\text{Cramer's V} = \sqrt{\frac{\chi^2}{n \min(r-1, c-1)}}$$

- Pearson contingency coef.

$$\text{Pearson C} = \sqrt{\frac{\chi^2}{\chi^2 + n}}$$

CMH tests for ordinal factors

Three types of CMH tests:

Non-zero correlation

- Use when *both* row and column variables are ordinal.
- CMH $\chi^2 = (N-1)r^2$, assigning scores (1, 2, 3, ...)
- most powerful for *linear* association

Row/Col Mean Scores Differ

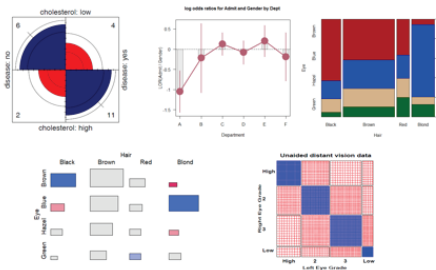
- Use when *only one* variable is ordinal
- Analogous to the Kruskal-Wallis non-parametric test (ANOVA on rank scores)

General Association

- Use when *both* row and column variables are nominal.
- Similar to overall Pearson χ^2 and Likelihood Ratio G^2 .

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Visualizing association



Observer agreement

- **Inter-observer agreement** often used as a way to assess reliability of a subjective classification or assessment procedure
 - → square table, Rater 1 × Rater 2
 - Levels: diagnostic categories (normal, mildly impaired, severely impaired)
- **Agreement vs. Association:** Ratings can be strongly associated without strong agreement
- **Marginal homogeneity:** Different frequencies of category use by raters affects measures of agreement
- **Measures of Agreement:**
 - Intraclass correlation: ANOVA framework—multiple raters!
 - Cohen's κ : compares the observed agreement, $P_o = \sum p_{ii}$, to agreement expected by chance if the two observer's ratings were independent, $P_e = \sum p_{i.} p_{.i}$.

$$\kappa = \frac{P_o - P_e}{1 - P_e}$$

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04: Loglinear models, mosaic displays

- Mosaic plots use sequential splits to show marginal and conditional frequencies in an n -way table
 - Shading: **sign** and **magnitude** of residuals → contributions to χ^2
 - Shows the pattern of association not accounted for
 - Permuting rows/cols often helps
- Loglinear models
 - Express associations with ANOVA-like interaction terms: $A*B, A*C$
 - Joint independence: $[AB][C] \equiv A * B + C$
 - Conditional independence: $[AC][BC] \equiv A \perp B | C$
 - Fitting models \equiv “cleaning the mosaic”
 - Response models: include all associations among predictors
- Sequential / partial plots & models
 - Sequential: Decompose all associations: $V_1; V_2 | V_1; V_3 | \{V_1, V_2\}, \dots$
 - Partial: Decompose conditional associations: $[V_1, V_2] | V_3 = \{a, b, \dots\}$

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Loglinear models: Perspectives

Loglinear models grew up and developed from three different ideas and ways of thinking about notions of independence in frequency data

- **Loglinear approach:** analog of ANOVA; associations are ~ interactions
- **glm() approach:** analog of general regression model, for $\log(\text{Freq})$, with **Poisson** dist^0 of errors
- **Logit models:** Loglinear, simplified for a **binary** response

Reduced models

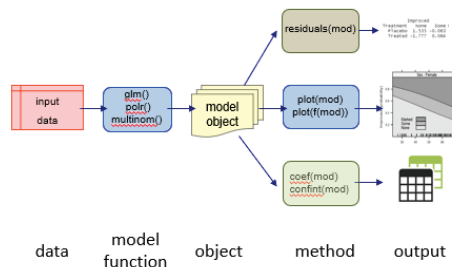
- For a three-way table there is a range of models between mutual independence, $[A][B][C]$, and the saturated model, $[ABC]$
- Each model has an independence interpretation:
 - $[A][B] \equiv A \perp B \equiv A$ independent of B
- Special names for various **submodels**

Table: Log-linear Models for Three-Way Tables

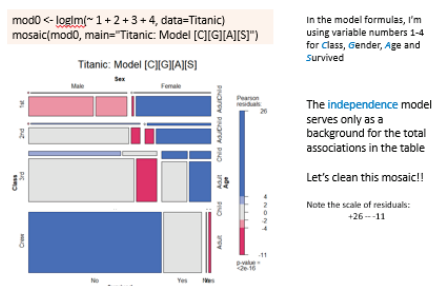
Model	Model symbol	Interpretation
Mutual independence	$[A][B][C]$	$A \perp B \perp C$
Joint independence	$[AB][C]$	$(AB) \perp C$
Conditional independence	$[AC][BC]$	$(A \perp B) C$
All two-way associations	$[AB][AC][BC]$	homogeneous assoc.
Saturated model	$[ABC]$	ABC interaction

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Model-based methods: Fitting & graphing



Fitting & visualizing models



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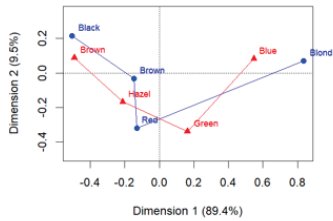
05: Correspondence analysis

- CA is an exploratory method designed to account for association (Pearson χ^2) in a small number of dimensions
 - Row and column scores provide an **optimal scaling** of the category levels
 - Plots of these can suggest an explanation for association
- CA uses the **singular value decomposition** to approximate the matrix of residuals from independence
- Standard and principal coordinates have different geometric properties, but are essentially re-scalings of each other
- Multi-way tables can be handled by:
 - Stacking approach—collapse some dimensions interactively to a 2-way table
 - Each way of stacking → a loglinear model
 - MCA analyzes the full n -way table using an indicator matrix or the **Burt** matrix

Given a new 2-way table, my first thought is nearly always: `plot(ca(mytable))`

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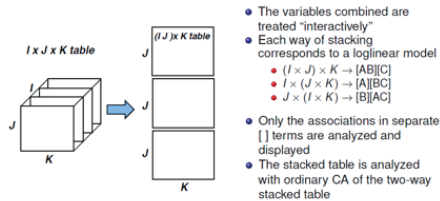
plot(haireye.ca, lines=TRUE)



- Rough interpretation: row/col points "near" each other are positively associated (independence residuals $d_{ij} >> 0$)
- Dim 1: 89.4% of χ^2 (dark \rightarrow light)
- Dim 2: 9.5% of χ^2 (Red/Green vs. others)

Multi-way tables: Stacking

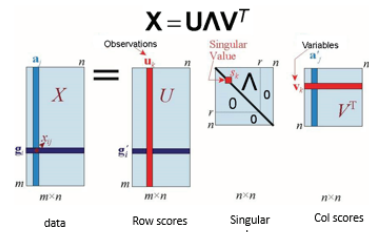
A 3-way table of size $I \times J \times K$ can be sliced and stacked as a two-way table in several ways



- The variables combined are treated "interactively"
- Each way of stacking corresponds to a loglinear model
 - $(I \times J) \times K \rightarrow [AB][C]$
 - $I \times (J \times K) \rightarrow [A][BC]$
 - $J \times (I \times K) \rightarrow [B][AC]$
- Only the associations in separate $[]$ terms are analyzed and displayed
- The stacked table is analyzed with ordinary CA of the two-way stacked table

Singular value decomposition

The singular value decomposition (SVD) is a basic technique for factoring a matrix and for matrix approximation
For an $m \times n$ matrix X of rank $r \leq \min(m, n)$ the SVD of X is:



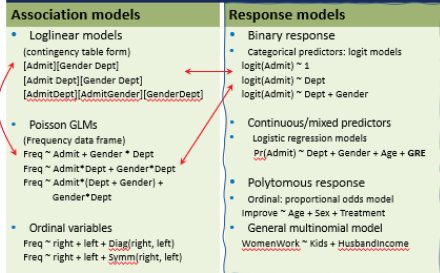
Multiple correspondence analysis

- Extends CA to n -way tables
- Useful when simpler stacking approach doesn't work well, e.g., 10 categorical attitude items
- Analyzes all pairwise bivariate associations. Analogous to:
 - Correlation matrix (numbers)
 - Scatterplot matrix (graphs)
 - All pairwise χ^2 tests (numbers)
 - Mosaic matrix (graphs)
- Provides an optimal scaling of the category scores for each variable
- Can plot all factors in a single plot
- An extension, *joint correspondence analysis*, gives a better account of inertia for each dimension

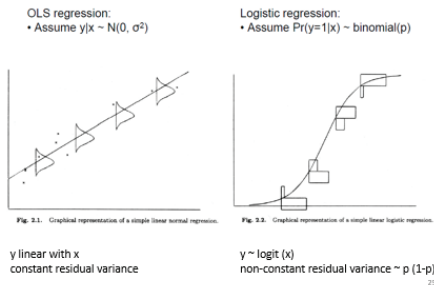
06: Logistic regression

- loglm() provides only overall tests of model fit
- Model-based methods, glm(), provide hypothesis tests, CIs & tests for individual terms
- Logistic regression: A glm() for a binary response
 - linear model for the log odds $\Pr(Y=1)$
 - All similar to classical ANOVA, regression models
- Plotting
 - Conditional, full-model plots show data and fits
 - Effect plots show predicted effects averaged over others
- Model diagnostics
 - Influence plots are often informative

Modeling approaches: Overview

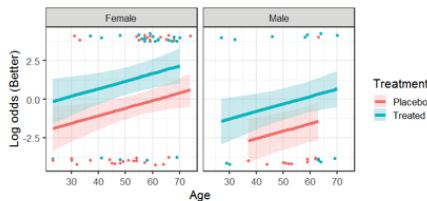


Linear regression vs Logistic regression



Full-model plot

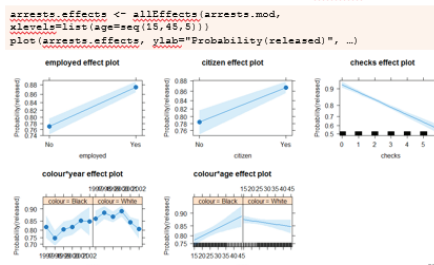
Plotting on the logit scale shows the additive effects of age, treatment and sex
NB: easier to compare the treatment groups within the same panel



These plots show model uncertainty (confidence bands)
Littered points show the data

Effect plots: allEffects

All high-order terms can be viewed together using plot(allEffects(mod))



07: Logistic regression: Extensions

- Polytomous responses
 - m response categories $\rightarrow (m-1)$ comparisons (logits)
 - Different models for ordered vs. unordered categories
- Proportional odds model
 - Simplest approach for ordered categories
 - Assumes same slopes for all logits
 - Fit with MASS::polr()
 - Test PO assumption with VGAM::vglm()
- Nested dichotomies
 - Applies to ordered or unordered categories
 - Fit $m - 1$ separate independent models \rightarrow Additive G^2 values
- Multinomial logistic regression
 - Fit $m - 1$ logits as a single model
 - Results usually comparable to nested dichotomies, but diff interpretation
 - R: nnet::multinom()

Exploratory plots

Before fitting models, it is useful to explore the data with conditional **ggplots**.

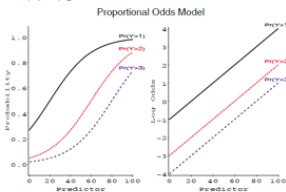


- Consider a logistic regression model for each logit:

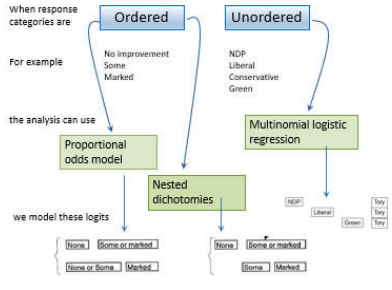
$$\text{logit}(\theta_{f1}) = \alpha_1 + \mathbf{X}_f^T \beta_1 \quad \text{None vs. Some/Marked}$$

$$\text{logit}(\theta_{f2}) = \alpha_2 + \mathbf{X}_f^T \beta_2 \quad \text{None/Some vs. Marked}$$

- Proportional odds assumption: regression functions are parallel on the logit scale i.e., $\beta_1 = \beta_2$.



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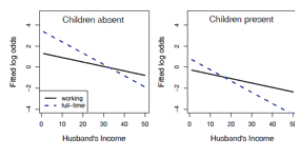
Nested dichotomies: Interpretation

Write out the predictions for the two logits, and compare coefficients:

$$\log \left(\frac{\text{Pr}(\text{working})}{\text{Pr}(\text{not working})} \right) = 1.336 - 0.042 \text{H\$} - 1.576 \text{kids}$$

$$\log \left(\frac{\text{Pr}(\text{fulltime})}{\text{Pr}(\text{parttime})} \right) = 3.478 - 0.107 \text{H\$} - 2.652 \text{kids}$$

Better yet, plot the predicted log odds for these equations:



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08: Extending loglinear models

- Loglinear models, as originally formulated, were quite general, but treated all table variables as **unordered** factors
 - The GLM perspective is more general, allowing quantitative predictors and handling **ordinal factors**
 - The logit model give a simplified approach when one variable is a **response**
- Models for **ordered factors** give more powerful & focused tests
 - $L \times L$, R , C and $R+C$ models **assign scores** to the factors
 - $RC(1)$ and $RC(2)$ models **estimate** the scores from the data
- Models for **square tables** allow testing structured questions
 - Quasi-independence: ignoring diagonals
 - symmetry & quasi-symmetry
 - theory-specific "topological" models
- These methods can be readily combined to analyze complex tables

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Logit models

For a **binary** response, each loglinear model is equivalent to a logit model (logistic regression, with categorical predictors)

- e.g., $\text{Admit} \perp \text{Gender} \mid \text{Dept}$ (conditional independence = $[\text{AD}][\text{DG}]$)

$$\log m_{ijk} = \mu + \lambda_i^A + \lambda_j^D + \lambda_k^G + \lambda_{ij}^{AD} + \lambda_{jk}^{DG}$$

So, for admitted ($i = 1$) and rejected ($i = 2$), we have:

$$\log m_{1jk} = \mu + \lambda_1^A + \lambda_j^D + \lambda_k^G + \lambda_{1j}^{AD} + \lambda_{jk}^{DG} \quad (1)$$

$$\log m_{2jk} = \mu + \lambda_2^A + \lambda_j^D + \lambda_k^G + \lambda_{2j}^{AD} + \lambda_{jk}^{DG} \quad (2)$$

Thus, subtracting (1)-(2), terms not involving Admit will cancel:

$$L_{jk} = \log m_{1jk} - \log m_{2jk} = \log(m_{1jk}/m_{2jk}) = \log \text{odds of admission}$$

$$= (\lambda_1^A - \lambda_2^A) + (\lambda_j^{AD} - \lambda_j^{DG})$$

$$= \alpha + \beta_j^{\text{Dept}} \quad (\text{renaming terms})$$

where, α : overall log odds of admission; β_j^{Dept} : effect on admissions of department

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Models for ordered categories

Consider an $R \times C$ table having **ordered** categories

- In many cases, the RC association may be described more simply by assigning numeric scores to the row & column categories.
- For simplicity, we consider only integer scores, $1, 2, \dots$ here
- These models are easily extended to stratified tables

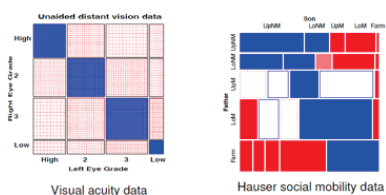
R:C model	μ_{ij}^{RC}	df	Formula
Uniform association	$i \times j \times \gamma$	1	$i: j$
Row effects	$a_i \times j$	$(I - 1)$	$R: j$
Col effects	$i \times b_j$	$(J - 1)$	$i: C$
Row+Col eff	$ja_i + ib_j$	$I + J - 3$	$R: j + i: C$
RC(1)	$\phi_{ij} \times \gamma$	$I + J - 3$	Mult (R, C)
Unstructured (R:C)	μ_{ij}^{RC}	$(I - 1)(J - 1)$	$R: C$

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Square tables

Square tables arise when the row and column variables have the **same** categories, often **ordered**

Special loglinear models allow us to tease apart different **reasons** for association

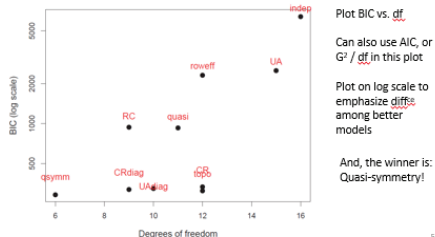


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Model comparison plots

When there are more than a few models, a **model comparison plot** can show the trade-off between goodness-of-fit and parsimony

- This sorts the models by both **fit** & **complexity**



Plot BIC vs. **df**

Can also use AIC, or G^2 / df in this plot

Plot on log scale to emphasize **diff:** among better models

And, the winner is: Quasi-symmetry!

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09: GLMs for Count Data

- GLMs provide a unified framework for linear models
 - Different families, all estimated in the same way
 - link function and associated variance function
- For count data, starting from $\log(\mu) = \mathbf{X} \beta$, $\mu | \mathbf{X} \sim \text{Poisson}$:
 - Overdispersion → quasi-poisson, negative binomial
 - Standard tools for assessing model fit
- Excess zero counts introduce new ideas & methods
 - ZIP model: structural model for the 0s
 - Hurdle model: random model for 0s, 2nd model for $Y > 0$
- In all this, we rely on data & model **plots** for understanding

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Canonical links and variance functions

- For every distribution family, there is a default, **canonical link** function
- Each one also specifies the expected relation between the mean and **variance**

Table 11.2: Common distributions in the exponential family used with generalized linear models and their canonical link and variance functions

Family	Notation	Canonical link	Range of η	Variance function, $V(\mu \eta)$
Gaussian	$N(\mu, \sigma^2)$	Identity: μ	$(-\infty, +\infty)$	σ^2
Poisson	$Poi(\mu)$	$\log_e(\mu)$	$0, 1, \dots, \infty$	μ
Negative-Binomial	$NBin(\mu, \theta)$	$\log_e(\mu)$	$0, 1, \dots, \infty$	$\mu + \mu^2/\theta$
Binomial	$Bin(n, \mu)/n$	$\log_e(\mu)$	$0, 1, \dots, n/n$	$\mu(1 - \mu)/n$
Gamma	$G(\mu, \nu)$	μ^{-1}	$(0, +\infty)$	μ^2/ν
Inverse-Gaussian	$I(G(\mu, \nu))$	μ^{-2}	$(0, +\infty)$	μ^3/ν^2

Choose a basic family:

- Get a default, canonical link, $g(\mu)$
- Also get a variance function for free!

Quasi-poisson models

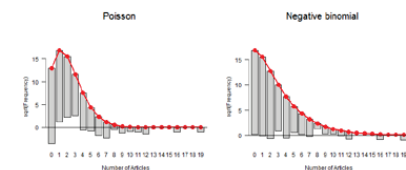
- The quasi-poisson model allows the dispersion, ϕ , to be a free parameter, estimates with other coefficients
- The conditional variance is allowed to be a multiple of the mean

$$\text{Var}(y_i | \eta_i) = \phi \mu_i$$

- This model is fit with `glm()` using `family=quasipoisson`
 - The estimated coefficients $\hat{\beta}$ are **unchanged**
 - The standard errors are multiplied by ϕ^2
 - Peace, order & good government is restored!

First, look at rootgrams:

```
plot(goodfit(PhdPubsArticles), xlab = "Number of Articles",
     main = "Poisson")
plot(goodfit(PhdPubsArticles, type = "binomial"),
     xlab = "Number of Articles", main = "Negative binomial")
```



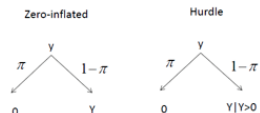
One reason the Poisson doesn't fit: excess 0s (some never published?)

Q: What might some other reasons be?
Think back to assumptions: independent obs; constant probs; unmodelled vars

Models for excess zeros

Two types of models, with different mechanisms for zero counts

- zero-inflated models:** The responses with $y_i = 0$ arise from a mixture of structural, always 0 values, with $\Pr(y_i = 0) = \pi$ and the rest, which are random 0s, with $\Pr(y_i = 0) = 1 - \pi$
- hurdle models:** One process determines whether $y_i = 0$ with $\Pr(y_i = 0) = \pi$. A second process determines the distribution of values of positive counts, $\Pr(y_i | y_i > 0)$



10: Models for log odds & LORs

- Logit models for a binary response generalize readily to a polytomous response
 - Models for log odds, familiar interpretation
 - Handles 3+ way table, ordinal variables
 - Simple plots for interpretation
- Generalized odds ratios handle bivariate responses
 - Simple linear models for LOR
 - Easy to model log odds for each response and the LOR simultaneously
 - Easy to visualize results

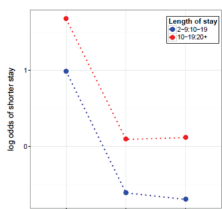
Main ideas

- Familiar case— Binary responses:
 - Every loglinear model for a binary response has an equivalent form in terms of **log odds** ["logit" models]
 - Log odds models have simple interpretations
 - Data + model plots give simple descriptions of data and models
- Extend to two-way ($I \times J$) and three-way ($I \times J \times K_1 \dots$) tables:
 - Log odds as **contrasts** in $\log(n)$
 - Variety of simple models for log odds (ANOVA-like)
 - Easily incorporate **ordinal** variables
 - Data + model plots give simple descriptions of data and models
- Generalized log odds ratios capture associations between two **focal variables**
 - Simple linear models for LOR
 - Direct visualization (Data + model plots) \implies more sensitive comparisons

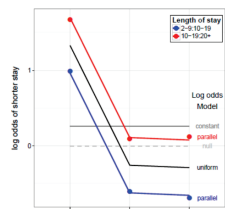
Based on my CARME (2015) presentation, <https://www.datavis.ca/papers/CARME2015-2x2.pdf>

Visualizing log odds and models

Plots of observed and fitted log odds: easy interpretation of data and models



Data plot: Observed log odds



Data + Model plot (fitted log odds)

Models for log odds

A variety of simple models can be specified in terms of log odds:

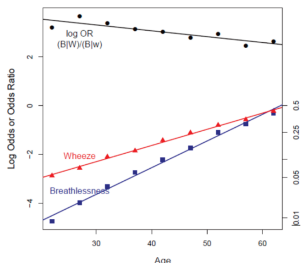
Table: Models for adjacent log odds in an $I \times J$ table with B as the response

Model	log odds parameters	degrees of freedom
null log odds	$\psi_{i,j}^{AB} = 0$	$I(J-1)$
constant log odds	$\psi_{i,j}^{AB} = \psi$	$I(J-1) - 1$
uniform B log odds	$\psi_{i,j}^{AB} = \psi^A$	$I(J-2)$
parallel log odds	$\psi_{i,j}^{AB} = \psi_i^A + \psi_j^B$	$(I-1)(J-2)$
saturated	$\psi_{i,j}^{AB}$ unspecified	

- The log odds, $\psi_{i,j}^{AB}$ can be viewed as entries in an $I \times (J-1)$ table
- These models are analogous to ANOVA tests of the A, B and A + B effects in this table.

These are similar to some `glm()` models we looked at before, but here in a simpler way by thinking in terms of log odds

Linear model for log odds and log odds ratios



Log odds & LORs have similar scales, so it is not terrible to plot them together

Going further: What did I leave out?

- This course, like DDAR, takes an entirely frequentist approach
 - Treats model parameters as **fixed**, constants
 - p -values, CIs interpreted as “over many repeated samples ...” --- 95% would contain the true value
 - Never really know if your sample gives correct inferences
- Bayesian approach
 - Treat parameters as random variables, incorporating **prior beliefs** and updating them with observed data to produce a **posterior distribution**. Much more intuitive.
 - A Bayesian 95% **credible interval** **directly** states there is a 95% probability the parameter lies within that range

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Fisher vs. Bayes



- Small samples, sparse data
 - Frequentist methods may fail to converge, or require fixups (add 0.5 to avoid 0 cells) for χ^2 tests to be valid
 - Bayesian approaches use priors to “smooth” empty cells, \Rightarrow more robust estimates in high-dimensional tables
- Logistic regression
 - Frequentist: relies on large-sample asymptotics for inference
 - Bayesian: provides full posterior distⁿ for odds ratios, β s
- Computation
 - Frequentist: very fast, particularly for large datasets
 - Bayesian:
 - Requires Bayesian sampling, e.g., MCMC; did my chains converge??
 - Originally (in DDAR): software quite limited, awkward to use
 - Now, much easier to use.

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Bayesian advances

- Regularization methods: stabilize inference with small or zero counts
- New variable selection methods
 - Effect clustering: merge levels of categorical predictor that have the same effect \Rightarrow more stable in high-D problems
- Applications:
 - Clinical trials: Bayesian adaptive designs can use historical data to speed up decision on efficacy of drug or treatment
 - Epidemiology: Bayesian hierarchical models used in disease mapping (estimating risk across areas), pooling strength

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JAMA[®]

Published online March 23, 2026

PERSPECTIVE

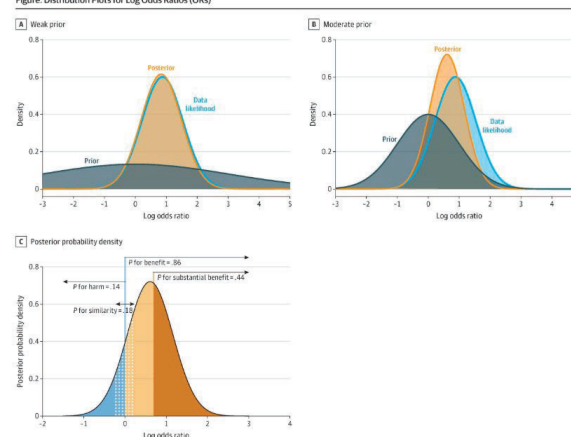
BREAKING NEWS: Bayesian methods adopted in JAMA recommendations to FDA

<https://t.co/7mMecV2E11>

Embracing Bayesian Methods in Clinical Trials
FDA's Long-Awaited Draft Guidance

J. Jack Lee, PhD; Frank E. Harrell Jr, PhD; Lisa M. LaVange, PhD; David J. Spiegelhalter, PhD

Figure. Distribution Plots for Log Odds Ratios (ORs)



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Hierarchical, longitudinal data

- All methods in this course — χ^2 tests, loglinear models, logistic regression, ...— assume observations are **independent**
- Fails when:
 - **same person** is measured at multiple time points (longitudinal / repeated measures), or
 - people are **nested within groups** such as schools, clinics, or geographic regions (hierarchical).
- Ignoring dependence:
 - Underestimated std. errors: CIs too narrow, p -values too small (🤩 phantom precision)
 - Biased parameter estimates

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Longitudinal CDA designs

Data structure: n subjects, each measured at T occasions.

- Categorical response: — e.g., disease present/absent, symptom severity (ordered), count of events in a period.
- Examples:
 - Panel surveys: political preference over election cycles
 - Clinical: patients worse, stable, improved at 4, 8, 12 weeks
 - Epidemiology: hospital visits per patient over 3 years

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Hierarchical (clustered / multilevel) designs

Data structure: Individuals at level 1 are nested within groups at level 2 (maybe: level 3, ...).

- The grouping creates dependence: two students in same school, or two patients at same clinic, share unmeasured **contextual influences**.
- Examples:
 - Students (level 1) nested in classrooms (level 2) nested in schools (level 3); binary outcome = passed a standardized test.
 - Patients nested in hospitals; ordinal outcome = satisfaction rating.
- Why standard methods fail:
 - Observations within the same cluster are positively correlated (intraclass correlation, ICC)
 - Std. logistic / Poisson regression treats all as independent, inflating the effective sample size, giving over-confident inference
 - Group-level predictors tested against wrong error term.

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Modeling approaches: GEE

- Marginal models via Generalized Estimating Equations (GEE)
 - Fit a population-average GLM (logistic, Poisson, ordinal, ...) as usual, but explicitly model the within-S correlation structure (exchangeable, AR(1), unstructured, ...).
 - GEE adjusts standard errors so inference is valid even if the correlation structure is mis-specified (robust/“sandwich” SEs).
 - Interpretation: coefficients describe the **average effect in the population**, not effects for a specific individual.
- R: `geepack::geeglm()`, `gee::gee()`

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General Linear Mixed Models (GLMM)

The natural extension of mixed models to non-normal outcomes. For a binary response with two-level nesting:

$$\text{logit}[P(Y_{ij} = 1)] = \mathbf{x}_{ij}^T \boldsymbol{\beta} + u_j$$

where $u_j \sim N(0, \sigma_u^2)$ is a random intercept for group j .

- $\boldsymbol{\beta}$: fixed effects (individual- and group-level predictors).
- σ_u^2 : variance of group-level random effects
- Can add **random slopes** if effect of predictor varies across groups
- R: `lme4::glmer()`, `glmmTMB::glmmTMB()`, `brms::brm()` (Bayesian)

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Transition (Markov) models

- Model response at time t as a function of covariates *and* the lagged response $Y_{i,t-1}$ (or several lags).
- Directly captures how **prior** state predicts current state
- Natural for panel /clinical data where change over time is the scientific focus.
- Equivalent to a conditional logistic regression on each transition.

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Connections to course topics

Course topic	Longitudinal/hierarchical extension
Logistic regression (binary)	GLMM with logit link; GEE with binary family
Ordinal logistic (proportional odds)	Mixed proportional-odds model (<code>ordinal::clmm()</code>)
Poisson / negative binomial (counts)	Poisson GLMM; NB GLMM (<code>glmmTMB</code>)
Multinomial logistic	Mixed multinomial (<code>mclogit::mblogit()</code>)
Loglinear models	Mixed log-linear for repeated cross-classifications

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R software summary

Package	Primary use
<code>lme4</code>	GLMMs: <code>glmer()</code> for binary/count; fast, widely used
<code>glmmTMB</code>	GLMMs with more families (NB, zero-inflated, ordinal), random slopes
<code>geepack</code>	GEE: <code>geeglm()</code> ; marginal models, robust SEs
<code>ordinal</code>	Mixed ordinal models: <code>clmm()</code>
<code>brms</code>	Bayesian GLMMs via Stan; full posterior for all variance components

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WHAT! There's even more??

Five other areas in a totally comprehensive CDA course:

1. Latent Class Analysis (LCA)
2. Survival Analysis / Event History Models
3. Item Response Theory (IRT)
4. Missing Data & Multiple Imputation
5. Regularized Logistic Regression

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Latent Class Analysis & Survival Analysis

Latent Class Analysis (LCA)

- Factor analysis for categorical indicators: finds latent subgroups explaining associations among observed variables
- Widely used for diagnostic categories, behavioral typologies, health risk profiles
- R: `poLCA`, `mclust`
- Agresti Ch. 13; Collins & Lanza (2010), *Latent Class and Latent Transition Analysis*

Survival Analysis / Event History Models

- Outcome is *time to event* (relapse, death, dropout, recovery) — possibly censored
- Workhorses: discrete-time logistic regression; Cox proportional hazards
- Very common in clinical and health psychology
- R: `survival`, `flexsurv`; Agresti Ch. 12; Hosmer, Lemeshow & May (2008)

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Item Response Theory & Missing Data

Item Response Theory (IRT)

- Models $P(\text{correct or endorsed response})$ as a function of a latent trait (ability, attitude)
- Rasch, 2PL, 3PL models are logistic regression with a latent predictor
- Natural extension of GLMs — directly relevant to psychometrics
- R: `ltm`, `mirt`, `TAM`; Embretson & Reise (2000)

Missing Data / Multiple Imputation

- Cannot mean-impute a categorical variable; complete-case analysis distorts association estimates
- MICE (multiple imputation via chained equations) handles categorical predictors and outcomes directly
- R: `mice`; van Buuren (2018), *Flexible Imputation of Missing Data* (free online)

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Regularized Logistic Regression

The problem: when ordinary logistic regression fails

- Many predictors → instability, perfect separation, near-collinearity
- LASSO (L_1) shrinks coefficients to zero → automatic variable selection
- Ridge (L_2) shrinks toward zero → stable estimation, all predictors retained
- Elastic net combines both penalties

Tools and further reading

- Frequentist complement to Bayesian regularization (shrinkage priors)
- R: `glmnet`
- Hastie, Tibshirani & Friedman (2009), *The Elements of Statistical Learning*, Ch. 4, 18

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CDA Dad Jokes

- **Q:** What did the χ^2 statistic say to her p -value?
 - **A:** You're the only one who can tell me if our relationship is real
- **Q:** Why was the mosaic plot so calm?
 - **A:** He had already made peace with all his residuals.
- **Q:** Why did logistic regression break up with his linear regression?
 - **A:** It's not you honey, it's me! My love is binary — I can only give you 0 or 1.
- A loglinear model walks into a bar.
 - Bartender says: "We don't serve your type here."
 - Model says: "Fine — I'll just fit a main effects model and assume no one here is associated."

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Your turn: Feedback?

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What did you like/dislike about 6136?

- Topics: what were the:
 - most interesting?
 - most boring?
 - most challenging/difficult?
- How useful were the web materials for the course?
 - Quizzes? R scripts? Resources / extra readings?
- What topics/activities did you learn most from?
- What gave you the most difficulty?
- How does this relate to your own work?

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Tips for next time ...

- What should I try to differently the next time?
 - More of X?
 - Less of Y?
 - Aspects of how the course is structured?
 - Evaluation?

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